

# Solutions of First Order Linear Partial Differential Equations Related to Urn Models and Central Limit Theorems

**(joint work with Michael Drmota)**

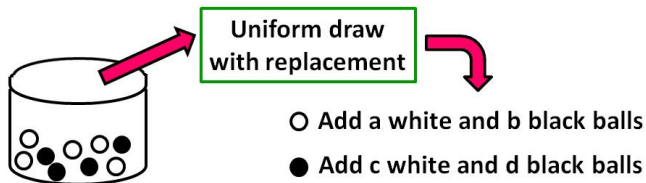
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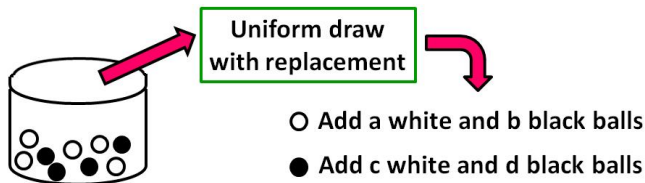
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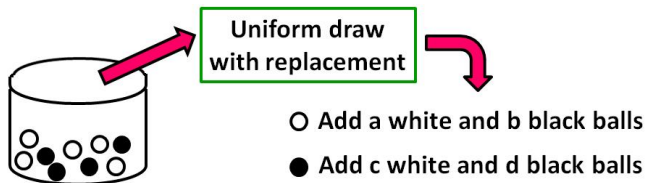
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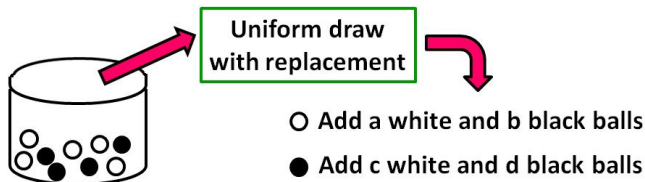
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- ▶ Rules for urn evolution  $\begin{pmatrix} a & b \\ c & d \end{pmatrix}$ ,  $a, b, c, d \in \mathbb{Z}$
- ▶ After a finite number of successive draws, we reach an absorbing state  $s = (j, k) \in \mathcal{S} = \mathbb{N} \times \mathbb{N}$  where the process of successive draws stops when the urn contains exactly  $j$  black and  $k$  white balls

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$$h_{n,m}(v) = \frac{n}{n+m} h_{n+a, m+c}(v) + \frac{m}{n+m} h_{n+c, m+d}(v)$$

## Three examples from Kuba and Panholzer (2007)

- ▶ The pill's problem has transition matrix  $\begin{pmatrix} -1 & 0 \\ 1 & -1 \end{pmatrix}$  and absorbing states  $\{(0, n) : n \geq 0\}$
- ▶ A variant of the pill's problem has transition matrix  $\begin{pmatrix} -1 & 0 \\ 1 & -2 \end{pmatrix}$  with absorbing states  $\{(0, n) : n \geq 0\} \cup \{(1, n) : n \geq 0\}$ .
- ▶ The cannibal urn has transition matrix  $\begin{pmatrix} -1 & 0 \\ 1 & -2 \end{pmatrix}$  with absorbing states  $\mathcal{S} = \{(0, n) : n \geq 0\} \cup \{(1, n) : n \geq 0\}$

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$$\frac{X_{n,m}}{\frac{n}{m} + \log m} \rightarrow \text{Exponential}(1) \quad (m \rightarrow \infty)$$

$$\frac{X_{n,m}}{n} \rightarrow \text{Beta}(1, m) \quad (\text{fixed } m \geq 1, n \rightarrow \infty)$$

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$$\begin{aligned} H(z, w) &= \frac{w}{v((1-z)^2 - w - ((v-1)/v))^2 (1-z - (v-1)/v)} \\ &+ \frac{(v-1)\sqrt{w}}{v^2((1-z)^2 - w - ((v-1)/v))^2} \\ &\times \arctan \left( \frac{\sqrt{w}\sqrt{(1-z)^2 - w - ((v-1)/v)^2}}{(1-z)^2 - w - (1-z)(v-1)/v} \right). \end{aligned}$$

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$$\frac{X_{n,2m}}{\frac{n}{\sqrt{m}} + 2\sqrt{m}} \rightarrow R \quad (m \rightarrow \infty \text{ } R \text{ has density } 2xe^{-x^2}, x \geq 0),$$

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$$\frac{X_{n,m} - \mathbb{E} X_{n,m}}{\sqrt{\text{Var} X_{n,m}}} \rightarrow N(0, 1) \quad (m+n \rightarrow \infty).$$

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First we have the system of first order ODE:

$$\frac{dz}{dt} = A(z, w), \quad \frac{dw}{dt} = B(z, w), \quad \frac{dH}{dt} = C(z, w)H + D(z, w; v),$$

where  $z = z(t)$ ,  $w = w(t)$ ,  $H = H(t)$  are functions in  $t$ .

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$$H(z, w; v) = \exp \left( \int_0^w \frac{C(f(Q_1(z, w), s), s)}{B(f(Q_1(z, w), s), s)} ds \right) \\ \times \left( \int_0^w \frac{D(f(Q_1(z, w), s), s; v)}{B(f(Q_1(z, w), s), s)} \exp \left( - \int_0^s \frac{C(f(Q_1(z, w), t), t)}{B(f(Q_1(z, w), t), t)} dt \right) ds \right)$$

## 2. Singularity analysis of $H(z, w; v)$

$$A(z, w)H_z + B(z, w)H_w - C(z, w)H = D(z, w; v) := \frac{a(z, w; v)}{(1 - b(z, w; v))^2}$$

$$\begin{aligned} H(z, w; v) &= \exp\left(\int_0^w \frac{C(f(Q_1(z, w), s), s)}{B(f(Q_1(z, w), s), s)} ds\right) \\ &\times \left(\int_0^w \frac{D(f(Q_1(z, w), s), s; v)}{B(f(Q_1(z, w), s), s)} \exp\left(-\int_0^s \frac{C(f(Q_1(z, w), t), t)}{B(f(Q_1(z, w), t), t)} dt\right) ds\right) \\ &:= \exp(K(z, w)) \int_0^w \frac{D(f(Q_1(z, w), s), s; v)}{B(f(Q_1(z, w), s), s)} \exp(-K(z, s)) ds \\ &= \int_0^w \frac{a(f(Q_1(z, w), s), s; v) \exp(-K(z, s)) / B(f(Q_1(z, w), s), s)}{(1 - b(f(Q_1(z, w), s), s; v))^2} ds \\ &:= \int_0^w \frac{L(z, w, s)}{(1 - b(f(Q_1(z, w), s), s; v))^2} ds \end{aligned}$$

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$$\forall z, w \in \mathbb{R}^+, \quad \frac{\partial f}{\partial s} = \frac{A(f, s)}{B(f, s)} < 0 \quad \Rightarrow \quad f(Q_1(z, w), 0) = 1$$
$$\Leftrightarrow z + w \rightarrow 1$$

## 2. Singularity analysis of $H(z, w; v)$ : when $v = 1$

By examples, we assume  $b(z, w; 1) = z$ . So

$$H(z, w; v) = \int_0^w \frac{L(z, w, s)}{(1 - f(Q_1(z, w), s))^2} ds$$

$$H(z, w; 1) = \sum_{n \geq 0, m \geq 1} \binom{n+m}{m} h_{n,m}(1) z^n w^m = \frac{1}{1-z-w} - \frac{1}{1-z}$$

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$H(z, w; 1)$  will not get singular if  $z + w \neq 1$ . But this representation for  $H(z, w; 1)$  holds if  $v$  is close to 1 and  $|1 - z - w| \geq \delta$  for some  $\delta > 0$ .



## 2. Singularity analysis of $H(z, w; v)$ : when $v$ close to 1

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$$H(z, w; v) \sim \frac{N(z, w; v)}{1 - b(f(Q_1(z, w), 0), 0; v)}$$

### 3. Asymptotics for bivariate generating functions

A generalized case in Pemantle and Wilson's book (2013)

**Lemma** Let  $f(z, w)$  be a bivariate generating function in the form

$$f(z, w) = \frac{N(z, w)}{D(z, w)},$$

where  $N$  and  $D$  are regular functions such that the system

$$D(z, w) = 0, \quad wD_w(z, w) = \rho z D_z(z, w)$$

has a unique positive and analytic solution  $z = z_0(\rho)$ ,  $w = w_0(\rho)$  for  $\rho \in [\alpha, \beta]$  such that

- ▶  $D_w(z_0(\rho), w_0(\rho)) \neq 0$
- ▶  $D(z, w) = 0$  has no other solutions for  $|z| \leq z_0(\rho)$ ,  $|w| \leq w_0(\rho)$
- ▶  $N(z_0(\rho), w_0(\rho)) \neq 0$

### 3. Asymptotics for bivariate generating functions (next)

Then we have uniformly for  $m/n \in [\alpha, \beta]$

$$[z^n w^m]f(z, w) \sim \frac{N(z_0(m/n), w_0(m/n))}{-z_0(m/n)w_0(m/n)D_z(z_0(m/n), w_0(m/n))} \cdot \frac{z_0(m/n)^{-n}w_0(m/n)^{-m}}{\sqrt{2\pi n\Delta(m/n)}},$$

where

$$\Delta(\rho) = \frac{D_{zz}D_w^2 - 2D_{zw}D_zD_w + D_{ww}D_z^2}{zD_z^3} + \frac{D_w^2}{z^2D_z^2} + \frac{D_w}{zwD_z} \Big|_{z=z_0(\rho), w=w_0(\rho)}$$

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- ▶ We fix the ratio  $m/n = \rho$  and  $h_\rho(w) := -\ln(z(w) + w^\rho)$

$$\begin{aligned} [w^{n\rho} z^n]f(z, w) &\sim \frac{1}{2\pi i} \int_{|w|=w_0(\rho)} \frac{N(z(w), w) z(w)^{-n} e^{nh(w)}}{-w \cdot z(w) D_z(z(w), w)} dw \\ &\sim \frac{N(z_0(\rho), w_0(\rho)) z_0(\rho)^{-n} w_0(\rho)^{-n\rho}}{-w_0(\rho) z_0(\rho) D_z(z_0(\rho), w_0(\rho)) \sqrt{2\pi n h''_\rho(w_0(\rho))}} \end{aligned}$$

1.  $w_0(\rho)$  is the saddle point of  $h_\rho(w)$
2.  $z_0(\rho) := z(w_0(\rho))$  and  $\Delta(\rho) := h''_\rho(w_0(\rho))$
3.  $h'_\rho(w) = 0 \Rightarrow wD_w(z, w) = \rho z D_z(z, w)$

## 4. Asymptotic probability generating function

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that is uniform in  $v$  (for  $v$  sufficiently close to 1) and also  $z_0(\rho; v)$  and  $w_0(\rho; v)$  denote the solutions of the system of equations

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If we fix the ratio  $\rho = m/n$  the leading asymptotics is

$$z_0(\rho; v)^{-n} w_0(\rho; v)^{-\rho n} = e^{h(\rho; v)n}$$

with  $h(\rho; v) = -\log z_0(\rho; v) - \rho \log w_0(\rho; v)$ .

## 5. Quasi-Power Theorem

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$$\mathbb{E}[v^{X_{\rho n, n}}] = \frac{[z^n w^{\rho n}]H(z, w; v)}{[z^n w^{\rho n}]H(z, w; 1)} \sim \frac{C(\rho; v)}{C(\rho; 1)} e^{(h(\rho; v) - h(\rho; 1))n}.$$

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$$\frac{X_{n, m} - \mathbb{E} X_{n, m}}{\sqrt{n}} \rightarrow \mathcal{N}(0, \sigma^2(m/n))$$

uniformly for  $m + n \rightarrow \infty$ ,  $m/n \in [\alpha, \beta]$ , where

$$\mathbb{E} X_{n, m} \sim \mu(m/n) n, \quad \text{Var} X_{n, m} \sim \sigma^2(m/n) n,$$

$$\mu(\rho) = \left. \frac{\partial}{\partial v} h(\rho; v) \right|_{v=1} > 0 \quad \text{and} \quad \sigma^2(\rho) = \left. \frac{\partial^2}{\partial v^2} h(\rho; v) \right|_{v=1} + \mu(\rho)$$



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# Thank You!